



Lévy-driven CARMA processes: Non-equidistant observations and local stationarity

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Introduction to Lévy-driven CARMA Processes



From ARMA to CARMA

Versatile class of auto-regressive moving-average processes

$$X_n - \varphi_1 X_{n-1} - \dots - \varphi_p X_{n-p} = \varepsilon_n + \theta_1 \varepsilon_{n-1} + \dots + \theta_q \varepsilon_{n-q}$$

Extensions to

- ▶ multivariate models (Vector ARMA)
- ▶ continuous-time models (CARMA)

Advantages of using time series models defined in continuous time:

- ▶ Allow handling of **irregularly spaced data** and **missing observations** (thus suitable for **high-frequency data**).
- ▶ Allows **consistent** estimation and inference at different **frequencies**

Main idea: Difference equations become differential equations

Problem: Definition, Properties and Estimation



CARMA process - definition

A second-order Lévy-driven continuous-time ARMA(p, q) process is defined (see e.g. Brockwell (2001), Brockwell (2009)) in terms of a *state-space representation* of the formal differential equation

$$a(D)Y(t) = b(D)DL(t), \quad t > 0. \quad (1)$$

where D denotes differentiation with respect to t , $(L(t))_{t \geq 0}$ is a Lévy process with $\sigma^2 = \mathbb{E}L(1)^2 < \infty$,

$$a(z) := z^p + a_1 z^{p-1} + \cdots + a_p,$$

$$b(z) := b_0 + b_1 z + \cdots + b_{q-1} z^{q-1} + b_q z^q,$$

with $b_q = 1$ and $q < p$.

The state-space representation

The **observation** and **state** equations are given by

$$Y(t) = \mathbf{b}^T \mathbf{X}(t) \quad (2)$$

$$d\mathbf{X}(t) = \mathbf{A}\mathbf{X}(t)dt + \mathbf{e}dL(t) \quad (3)$$

where

$$\mathbf{A} := \begin{bmatrix} 0 & 1 & \dots & 0 \\ 0 & 0 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & 1 \\ -a_p & -a_{p-1} & \dots & -a_1 \end{bmatrix}, \quad \mathbf{X}(t) := \begin{bmatrix} X(t) \\ X^{(1)}(t) \\ \vdots \\ X^{(p-2)}(t) \\ X^{(p-1)}(t) \end{bmatrix},$$

$$\mathbf{e} := [0, \dots, 0, 1]^T \text{ and } \mathbf{b} := [b_0, \dots, b_{p-1}]^T.$$

Lévy driven CARMA processes

Assumption

$$\mathbb{E}[L(1)] = 0 \text{ and } \mathbb{E}[L(1)]^2 = \sigma^2 < \infty.$$

The solution $\mathbf{X}(t)$ of (3) satisfies

$$\mathbf{X}(t) = e^{\mathbf{A}t} \mathbf{X}(0) + \int_0^t e^{\mathbf{A}(t-u)} \mathbf{e} dL(u). \quad (4)$$

If all eigenvalues of \mathbf{A} have strictly negative real parts, the process $(\mathbf{X}(t))_{t \in \mathbb{R}}$ given by

$$\mathbf{X}(t) = \int_{-\infty}^t e^{\mathbf{A}(t-u)} \mathbf{e} dL(u)$$

is a **causal strictly stationary solution** of (3) for $t \in \mathbb{R}$ with the corresponding CARMA process

$$Y(t) = \int_{-\infty}^t \mathbf{b}^T e^{\mathbf{A}(t-u)} \mathbf{e} dL(u). \quad (5)$$

ACF vs Spectral Density

- ▶ The spectral density function

$$f(\omega) := \int_{\mathbb{R}} r(h) e^{-i\omega h} dh$$

where $r(h) = \text{Cov}(Y(0), Y(h))$ denotes the autocovariance function of Y .

- ▶ It holds

$$r(h) = \frac{1}{2\pi} \int_{\mathbb{R}} f(\omega) e^{i\omega h} d\omega$$

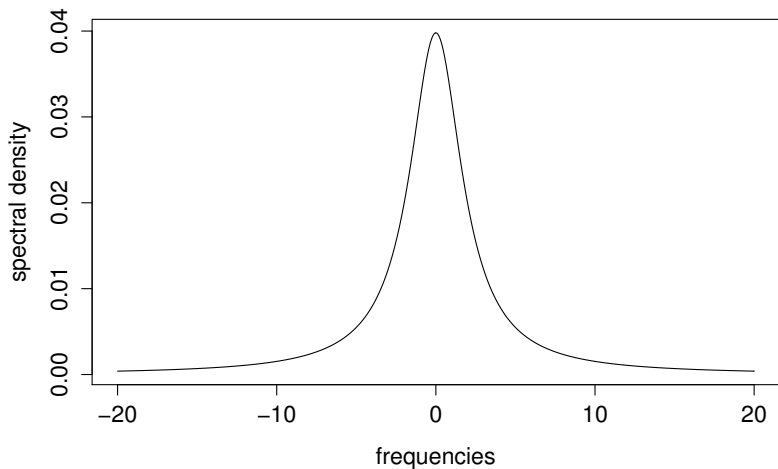
- ▶ For CARMA processes the spectral density function is of the form

$$f(\omega) = \frac{\sigma^2}{2\pi} \frac{|b(i\omega)|^2}{|a(i\omega)|^2}$$

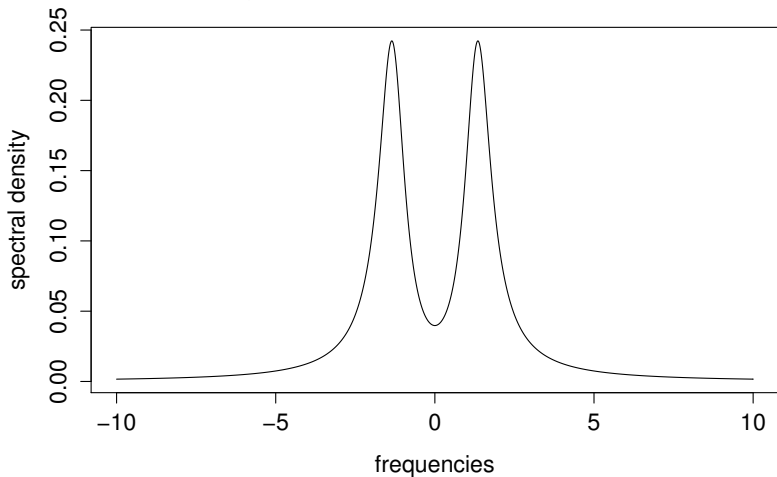
and is square integrable.

- ▶ The Fourier and inverse Fourier transform: a way of communicating between the time and frequency domain.

Spectral Density of CAR(1) Process (AR Coefficient=2)



Spectral Density of CARMA(2,1) Process ($[a_2, a_1, b_0] = [2, 1, 2]$)



Dependence structure

- ▶ **Markov properties:**

As a solution to a stochastic differential equation, the **state space representation \mathbf{X}** of a causal MCARMA process is a **strong Markov process**.

- ▶ **Mixing properties:**

For a causal MCARMA process the state space representation **\mathbf{X}** is **β -mixing** and **Y** is **strongly mixing**, both with **exponentially decaying mixing coefficients**.

In particular, **\mathbf{X}** and **Y** are **ergodic** (**\mathbf{X}** is also geometrically ergodic).

Sample path properties

- ▶ The sample paths of a (causal) CARMA(p, q) process Y with $p > q + 1$ are $(p - q - 2)$ -times differentiable and the $p - q - 2$ th derivative is absolutely continuous.
- ▶ If $p = q + 1$, then $\Delta Y(t) = b_q \Delta L(t)$.
- ▶ If the driving Lévy process L is a Brownian motion, then the sample paths of Y are continuous and $(p - q - 1)$ -times continuously differentiable, provided $p > q + 1$.



Statistics of Lévy-driven CARMA Processes: A Summary



What is known about estimating CARMA processes?

- ▶ Brockwell, Davis, and Yang (2011), Schlemm and St. (2012):
QML estimation of AR and MA parameters based on equidistant samples for (multivariate) Lévy-driven CARMA processes
(Anti-**Aliasing** Condition needed)
- ▶ Fasen and Kimmig (2016):
Order selection by information criteria for QML estimators
- ▶ Brockwell and Schlemm (2013):
reconstruction of the driving Lévy process using a high-frequency limit (for equidistant samples) and based on this GMM estimation of the parameters of the Lévy process (letting the observation time interval to ∞).
- ▶ Ferrazzano and Fuchs (2013):
Alternative approach to recovering the driving Lévy process for invertible CARMA processes under a high frequency limit

What is known about estimating CARMA processes?

- ▶ Fasen and Fuchs (2013b), Fasen and Fuchs (2013a):
Estimation of power transfer function (also for stable driving Lévy processes) from discrete-time periodograms under a high-frequency infinite time horizon limit using equidistant observations.
- ▶ Fasen (2014b), Fasen (2016):
(high-frequency infinite horizon) limiting results for sample autocovariances
- ▶ Gillberg (2006):
Estimation of non-equidistantly sampled (Brownian) CARMA processes by spline interpolation and using results for equidistant grids

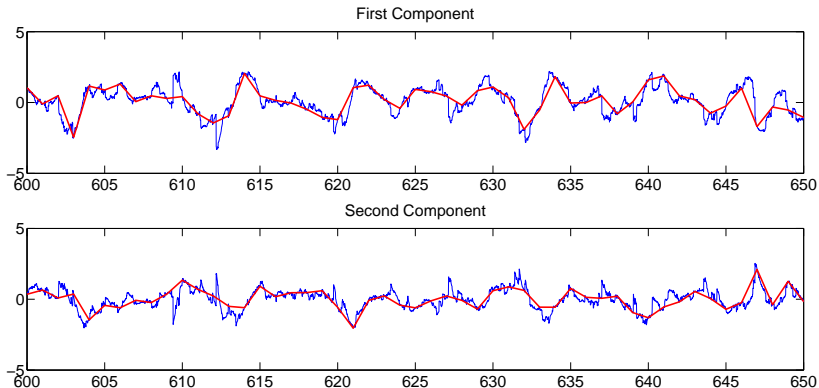
Estimation of Poisson Sampled CARMA Processes



The aliasing problem for equidistant observations

Assume we observe a (multivariate) CARMA process Y at discrete, equally spaced times

$$Y_n^{(h)} := Y(nh), \quad n \in \mathbb{Z}, \quad h > 0.$$



The aliasing problem for equidistant observations

- ▶ The autoregressive and moving average parameters and the variance of the Lévy process are **identifiable from the spectral density of Y** , if the autoregressive and moving average polynomial have no common zeros.
- ▶ The autoregressive and moving average parameters and the variance of the Lévy process are **NOT identifiable from the spectral density of $Y^{(h)}$** , if the autoregressive and moving average polynomial have no common zeros.
- ▶ **Anti-aliasing condition:** All eigenvalues of \mathbf{A} or equivalently all zeros of the autoregressive polynomial are in the set $\{z \in \mathbb{C} : -\pi/h < \Im z < \pi/h\}$.
- ▶ Under the anti-aliasing condition, the autoregressive and moving average parameters and the variance of the Lévy process are **identifiable from the spectral density of $Y^{(h)}$** , if the autoregressive and moving average polynomial have no common zeros.

Poisson Sampling

- ▶ Let N be a Poisson process with rate β and jump times $\{\tau_k\}$ which is independent of L .
- ▶ The CARMA process is observed at $\{\tau_k\}$ which is equivalent to observing the process Z given by

$$dZ(t) := Y(t) dN(t)$$

- ▶ Z can also be understood as an orthogonal random measure with spectral density

$$\phi_Z(\omega, \boldsymbol{\theta}) = \beta^2 \left[\frac{\sigma^2}{2\pi} \left| \frac{b(i\omega)}{a(i\omega)} \right|^2 + \frac{\sigma^2 \mathbf{b}^T \boldsymbol{\Sigma} \mathbf{b}}{2\pi\beta} \right] = \frac{\beta^2 \sigma^2}{2\pi} \left[\left| \frac{b(i\omega)}{a(i\omega)} \right|^2 + \frac{\mathbf{b}^T \boldsymbol{\Sigma} \mathbf{b}}{\beta} \right].$$

with $\boldsymbol{\theta} = (a_1, \dots, a_p, b_0, \dots, b_{q-1})$ and $\boldsymbol{\Sigma} = \int_0^\infty e^{Ay} \mathbf{e}_p \mathbf{e}_p^T e^{A^T y} dy$.

Identifiability under Poisson Sampling

Proposition (Lii and Masry (1992); Masry (1978))

*Assume the autoregressive and moving average polynomials have no common zeros. Then the autoregressive and moving average parameters and the variance of the Lévy process are **identifiable from the spectral density of Z** .*



A Whittle type Estimator

- ▶ We assume that we observe Z on $[0, T]$.

- ▶ Setting $g_1(\omega, \theta) := \left| \frac{b(i\omega)}{a(i\omega)} \right|^2 + \frac{\mathbf{b}^T \Sigma \mathbf{b}}{\beta}$ and

$$g_2(\theta) := \frac{\pi \mathbf{b}^T \Sigma \mathbf{b}}{\beta} + \int_{-\infty}^{\infty} \left(\frac{|b(ix)/a(ix)|^2}{1+x^2} \right) dx \text{ we define}$$

$$\hat{K}_T(\theta) = \frac{1}{2\pi T} \int_{-\infty}^{\infty} \left(\frac{\log g_1(\omega, \theta) - \log g_2(\theta)}{1 + \omega^2} \left| \sum_{k=1}^{N(T)} e^{-i\omega\tau_k} Y(\tau_k) \right|^2 \right) d\omega.$$

- ▶ Denoting the parameter space of the model by Θ the **Whittle-type estimator** is given by:

$$\hat{\theta}_T = \arg \max_{\theta \in \Theta} \hat{K}_T(\theta).$$

- ▶ θ_0 denotes the true autoregressive and moving average parameters.

Consistency

Theorem (Bosserhoff, Fechner and St. (2016))

Assume

- ▶ Θ is compact,
- ▶ the driving Lévy process L has finite fourth moments,
- ▶ for all $\theta \in \Theta$ the autoregressive and moving average polynomials have no common zeros.

Then

$$\hat{\theta}_T \xrightarrow[T \rightarrow \infty]{\mathbb{P}} \theta^0.$$

The Asymptotic Covariance Matrix

- ▶ Let $M \in \mathbb{R}^{(p+q) \times (p+q)}$ be defined by

$$M := W^{-1}QW,$$

where

- ▶ $W \in \mathbb{R}^{(p+q) \times (p+q)}$ is given by

$$w_{i,j} = - \int_{-\infty}^{\infty} \frac{\partial \log g(\omega, \theta^0)}{\partial \theta_i^0} \frac{\partial \log g(\omega, \theta^0)}{\partial \theta_j^0} \frac{\phi_Z(\omega, \theta^0)}{1 + \omega^2} d\omega,$$

- ▶ $Q \in \mathbb{R}^{(p+q) \times (p+q)}$ is given by

$$q_{i,j} = 2\pi \left\{ \int_{-\infty}^{\infty} \frac{\partial \log g(\omega, \theta^0)}{\partial \theta_i^0} \frac{\partial \log g(\omega, \theta^0)}{\partial \theta_j^0} \frac{\phi_Z^2(\omega, \theta^0)}{(1 + \omega^2)^2} d\omega \right. \\ \left. + \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{\partial \log g(\omega_1, \theta^0)}{\partial \theta_i^0} \frac{\partial \log g(\omega_2, \theta^0)}{\partial \theta_j^0} \frac{\phi_Z^{(4)}(\omega_1, -\omega_1, \omega_2, \theta^0)}{(1 + \omega_1^2)(1 + \omega_2^2)} d\omega_1 d\omega_2 \right\}.$$

- ▶ with $g(\omega, \theta) = g_1(\omega, \theta)/g_2(\theta)$

Asymptotic Normality

Theorem (Bosserhoff, Fechner and St. (2016))

Assume

- ▶ Θ is compact,
- ▶ the driving Lévy process L has finite moments of all orders,
- ▶ for all $\theta \in \Theta$ the autoregressive and moving average polynomials have no common zeros,
- ▶ and the matrix W is invertible.

Then $\sqrt{T}(\hat{\theta}_T - \theta^0)$ is asymptotically normally distributed with mean zero and covariance matrix M .

Both consistency and asymptotic normality are proven by verifying the conditions of Lii and Masry (1992) for Lévy-driven CARMA processes.

Estimating the Variance of the driving Lévy Process

Proposition (Bosserhoff, Fechner and St. (2016))

$$\hat{\sigma}_T^2 := \frac{2\pi}{\beta^2} \frac{\int_{-\infty}^{\infty} \frac{I_{Z,T}(\omega)}{1+\omega^2} d\omega}{\pi \frac{\hat{\mathbf{b}}^T \hat{\Sigma} \hat{\mathbf{b}}}{\beta} + \int_{-\infty}^{\infty} \left(\frac{|\hat{b}(i\omega)/\hat{a}(i\omega)|^2}{1+\omega^2} \right) d\omega}$$

is a consistent estimator of σ^2 .

Note that we throughout assumed the rate β to be known.

High-frequency irregularly sampled data



Our Approach

- ▶ Suppose that we observe a CARMA process at times $t_1 < t_2 < \dots < t_N$ which are not necessarily uniformly spaced on $[0, T]$. Let

$$h_{\max} := \max_{1 \leq i \leq N-1} (t_{i+1} - t_i)$$

be the maximal distance between observations.

- ▶ What can we say about the limiting behaviour of estimators when the length of the interval $T \rightarrow \infty$ and the grid size goes to zero (i.e. $h_{\max} \rightarrow 0$)?

Truncated Fourier Transform

The truncated continuous-time Fourier transform of the process Y at a frequency $\omega \in \mathbb{R}$:

$$\mathcal{F}_T(Y)(\omega) := \frac{1}{\sqrt{T}} \int_0^T Y(t) e^{-i\omega t} dt.$$

Lemma (Fechner and St. (2015))

Let Y be a Lévy-driven CARMA process. Then the *truncated Fourier transform* of the CARMA process Y at a fixed frequency $\omega \in \mathbb{R}$ is of the form

$$\begin{aligned} \mathcal{F}_T(Y)(\omega) &= \frac{1}{\sqrt{T}} \frac{b(i\omega)}{a(i\omega)} \int_0^T e^{-i\omega t} dL(t) \\ &\quad + \frac{1}{\sqrt{T}} \mathbf{b}^T (i\omega I - A)^{-1} (\mathbf{X}(0) - e^{-i\omega T} \mathbf{X}(T)) \end{aligned}$$

Truncated Fourier Transform: Covariances

Theorem (Fechner and St. (2015))

Let Y be a Lévy-driven CARMA process. Then for $\omega_1, \omega_2 \in \mathbb{R}$ we have

$$\mathbb{E}[\mathcal{F}_T(Y)(\omega_1)\mathcal{F}_T(Y)(\omega_2)] = \sigma^2 f_Y(\omega_1) + \frac{1}{T}K(T, \omega_1, \omega_2), \quad \omega_1 = -\omega_2 \quad (6)$$

and

$$\mathbb{E}[\mathcal{F}_T(Y)(\omega_1)\mathcal{F}_T(Y)(\omega_2)] = \frac{1}{T}K^1(T, \omega_1, \omega_2), \quad \omega_1 \neq -\omega_2, \quad (7)$$

where f_Y is the spectral density function of the process Y , K, K_1 are bounded functions of T .

Truncated Fourier Transform: Asymptotic Normality I

Proposition (Fechner and St. (2015))

Let

$$Z(T) := \frac{1}{\sqrt{T}} \frac{b(0)}{a(0)} \int_0^T dL(t).$$

Then

$$\text{dlim}_{T \rightarrow \infty} Z(T) = \text{dlim}_{T \rightarrow \infty} \mathcal{F}_T(Y)(0) \sim \mathcal{N} \left(0, \left(\frac{b(0)}{a(0)} \right)^2 \sigma^2 \right)$$

$$\text{dlim}_{T \rightarrow \infty} \left| \frac{a(0)Z(T)}{b(0)\sigma^2} \right|^2 = \text{dlim}_{T \rightarrow \infty} \left| \frac{a(0)\mathcal{F}_T(Y)(0)}{b(0)\sigma^2} \right|^2 \sim \chi^2(1)$$

Truncated Fourier Transform: Asymptotic Normality II

Proposition (Fechner and St. (2015))

Assume that $\omega \neq 0$. Put

$$\tilde{Z}(T) := \frac{1}{\sqrt{T}} \frac{b(i\omega)}{a(i\omega)} \int_0^T e^{-i\omega t} dL(t) \text{ and } Z(T) = \begin{bmatrix} \Re \tilde{Z}(T) \\ \Im \tilde{Z}(T) \end{bmatrix}.$$

Then

$$\text{dlim}_{T \rightarrow \infty} Z(T) = \text{dlim}_{T \rightarrow \infty} \begin{pmatrix} \Re \mathcal{F}_T(Y)(\omega) \\ \Im \mathcal{F}_T(Y)(\omega) \end{pmatrix} \sim \mathcal{N}(0, \Sigma),$$

where

$$\Sigma = \frac{\sigma^2}{2} \left| \frac{b(i\omega)}{a(i\omega)} \right|^2 \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}.$$

The result can be extended to the joint convergence at different frequencies which are asymptotically independent.

Non-equidistant grids: Set-Up

- ▶ Assume now that for a time horizon $T > 0$ we observe the CARMA process Y at $N^{(T)}$ non-equidistant time points

$$0 = x_0^{(T)} < x_1^{(T)} < \dots < x_{N^{(T)}-1}^{(T)} = T.$$

- ▶ Let $h_{\max}^{(T)} = \max_{j=0, \dots, N^{(T)}-2} (x_{j+1}^{(T)} - x_j^{(T)})$.
- ▶ We denote the **trapezoidal approximation** of

$$\mathcal{F}_T(Y)(\omega) := \frac{1}{\sqrt{T}} \int_0^T Y(t) e^{-i\omega t} dt$$

by

$$\mathcal{T}_T(Y)(\omega).$$

Non-equidistant grids: Error Bounds

Proposition (Fechner and St. (2015))

1. We have that

$$\|\mathcal{T}_T(Y)(\omega) - \mathcal{F}_T(Y)(\omega)\|_{L^2}^2 \leq \frac{C_1 C_2}{T} + C_1 N^{(T)2} h_{\max}^{(T)6}$$

where the constants C_1, C_2 can be chosen independent of T .

2. So if $\lim_{T \rightarrow \infty} N^{(T)} h_{\max}^{(T)3} = 0$, then $\mathcal{T}_T(Y)(\omega) - \mathcal{F}_T(Y)(\omega) \rightarrow 0$ as $T \rightarrow \infty$ in L^2 and thus in probability.

The proof relies on the smoothness of $e^{-i\omega t}$ and techniques adopted from Brockwell and Schlemm (2013) to non-equidistant grids and changing time horizons.

Non-equidistant grids: Asymptotic Normality I

Proposition (Fechner and St. (2015))

If $\lim_{T \rightarrow \infty} N^{(T)} h_{\max}^{(T)3} = 0$, then we have for the trapezoidal approximation of the truncated Fourier transform

$$\text{dlim}_{T \rightarrow \infty} \mathcal{T}_T(Y)(0) \sim \mathcal{N} \left(0, \left(\frac{b(0)}{a(0)} \right)^2 \sigma^2 \right)$$

$$\text{dlim}_{T \rightarrow \infty} \left| \frac{a(0)\mathcal{T}_T(Y)(0)}{b(0)\sigma^2} \right|^2 \sim \chi^2(1).$$

Non-equidistant grids: Asymptotic Normality II

Proposition (Fechner and St. (2015))

If $\lim_{T \rightarrow \infty} N^{(T)} h_{\max}^{(T)3} = 0$ and $\omega \neq 0$, then we have for the trapezoidal approximation of the truncated Fourier transform

$$\text{dlim}_{T \rightarrow \infty} \begin{pmatrix} \Re \mathcal{T}_T(Y)(\omega) \\ \Im \mathcal{T}_T(Y)(\omega) \end{pmatrix} \sim \mathcal{N}(0, \Sigma),$$

where where

$$\Sigma = \frac{\sigma^2}{2} \left| \frac{b(i\omega)}{a(i\omega)} \right|^2 \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}.$$

The result can again be extended to the joint convergence at different frequencies which are asymptotically independent.

Simulated Example I

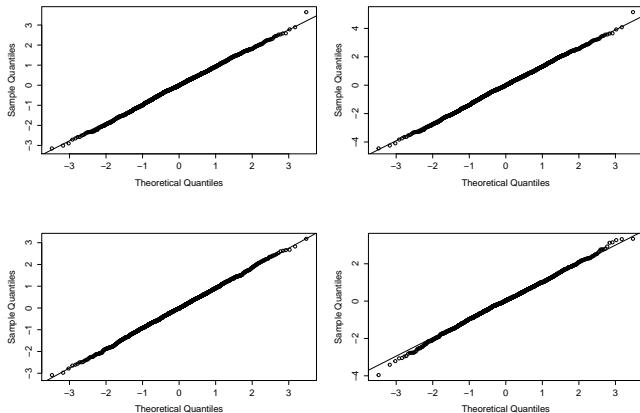


Figure : QQ plots for CAR(1) process driven by standard Brownian Motion noise with coefficients $a_1 = 1$ and $b_0 = 0$. Plotted are 2000 paths observed on a 2000 point (moderately) non-equidistant grid over the interval $[0, 1]$.

Simulated Example II

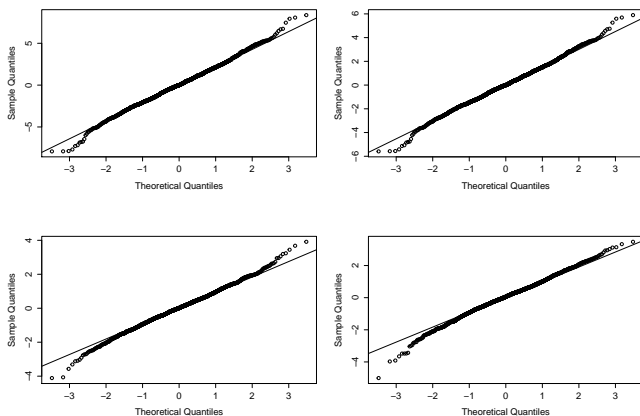


Figure : QQ plots for CARMA(2,1) process driven by a Variance Gamma noise with coefficients $a_2 = 1$, $a_1 = 2$, $b_1 = 1$. Plotted are 2000 paths observed on a 2000 point (moderately) non-equidistant grid over the interval $[0, 10]$

Locally stationary Lévy-driven CARMA Processes



Motivation

- ▶ Often observed time series behave stationary in the short run, but over longer periods the dynamics change.
- ▶ Hence *time-varying dynamics are encountered*.
- ▶ Various time-varying models have been defined in discrete time, e.g. *tvARMA*, *tvGARCH*.
- ▶ **Problem:** Statistical inference. How to define meaningful estimators and obtain appropriate asymptotics?
- ▶ Solution: “in-fill asymptotics” and consideration of sequences of processes
- ▶ Intensively studied in discrete time (see e.g. Dahlhaus (2012))

Stationary Moving Average Processes in Continuous Time

- ▶ Let L be a second order Lévy process.
- ▶ For $g \in L^2$ or $A \in L^2$ a linear/moving average process is given by

$$Y(t) = \int_{\mathbb{R}} g(t-u) dL(u) = \int_{\mathbb{R}} e^{i\mu t} A(\mu) \Phi_L(d\mu).$$

- ▶ Φ_L is the orthogonal random measure determined by

$$\Phi_L([a, b]) = \int_{\mathbb{R}} \frac{e^{-ita} - e^{-itb}}{2\pi it} dL(t).$$

- ▶ A and g are Fourier transforms of each other:

$$g(t-u) = \frac{1}{2\pi} \int_{\mathbb{R}} e^{i\mu(t-u)} A(\mu) d\mu \text{ and } A(\mu) = \int_{\mathbb{R}} e^{-i\mu(t-u)} g(t-u) du$$

Local Stationarity in Continuous Time

Definition (Bitter and St. (2016))

A sequence of linear stochastic processes $\{Y_N(t)\}_{N \in \mathbb{N}}$ given by

$$Y_N(t) = \int_{\mathbb{R}} g_N^0(Nt, Nt - u) dL(u) = \int_{\mathbb{R}} e^{i\mu t} A_N^0(Nt, \mu u) \Phi_L(d\mu)$$

is called **locally stationary** if one of the following equivalent conditions holds:

- ▶ there exists $g : \mathbb{R}^2 \rightarrow \mathbb{R}$ continuous in the first coordinate such that $g_N^0(Nt, \cdot) \rightarrow g(t, \cdot)$ for $N \rightarrow \infty$ in L^2 for all $t \in \mathbb{R}$.
- ▶ there exists $A : \mathbb{R}^2 \rightarrow \mathbb{C}$ continuous in the first coordinate such that $A_N^0(Nt, \cdot) \rightarrow A(t, \cdot)$ for $N \rightarrow \infty$ in L^2 for all $t \in \mathbb{R}$.

Local Stationarity: Immediate Properties

- ▶ For each $t \in \mathbb{R}$ $Y_N(t)$ converges in distribution to $\int_{\mathbb{R}} A(t, \mu) \Phi_L(d\mu) = \int_{\mathbb{R}} g(t, t-u) dL(u)$.
- ▶ For $t_1 \neq t_2$ $Y_N(t_1)$ and $Y_N(t_2)$ are asymptotically uncorrelated.



Locally Stationary OU type processes

Proposition (Bitter and St. (2016))

The sequence of stochastic processes $\{Y_N(t) : t \in \mathbb{R}\}_{N \in \mathbb{N}}$ defined by:

$$Y_N(t) = \int_{-\infty}^{Nt} e^{-\int_u^{Nt} a(\frac{s}{N}) ds} L(du), \quad (8)$$

is locally stationary, if the coefficient function $a : \mathbb{R} \mapsto \mathbb{R}$ satisfies:

(C1) $a(\cdot)$ is continuous on \mathbb{R} ,

(C2) $a(t) \geq \varepsilon$ for some $\varepsilon > 0$.

The limiting kernel / transition matrix is given by:

$$g(t, u) = \mathbf{1}_{\{u \geq 0\}} e^{-a(t)u} \quad \text{and} \quad A(t, \mu) = \frac{1}{a(t) - i\mu}.$$

Note: $dY(t) = a(t)Y(t)dt + L(dt)$ has the solution

$$Y(t) = \int_{-\infty}^t e^{-\int_u^t a(s) ds} L(du).$$

Wigner-Ville Spectrum

Definition

For a locally stationary process the Wigner-Ville spectrum for fixed $N \in \mathbb{N}$ is defined as

$$f_N(t, \lambda) := \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-i\lambda s} \text{Cov} \left(Y_N \left(t + \frac{s}{2N} \right), Y_N \left(t - \frac{s}{2N} \right) \right) ds$$

and the (time-varying) spectral density of the process $Y_N(t)$ as

$$f(t, \lambda) := \frac{\sum_L}{2\pi} |A(t, \lambda)|^2.$$

Convergence of the Wigner-Ville Spectrum

Theorem (Bitter and St. (2016))

Let $\{Y_N(t), t \in \mathbb{R}\}_{N \in \mathbb{N}}$, $Y_N(t) = \int_{\mathbb{R}} e^{i\mu Nt} A_N^0(Nt, \mu) \Phi_L(d\mu)$, be a locally stationary sequence of stochastic processes. Assume

- (i) For all $t, s \in \mathbb{R}$ $\|A_N^0(N(t \pm \frac{s}{2N}), \cdot) - A(t, \cdot)\|_{L^2(\mathbb{R}, \mathbb{C})} \xrightarrow{N \rightarrow \infty} 0$,
- (ii) A_N^0 and A are uniformly bounded in L^2 ,
- (iii) For all t, N, μ the derivatives of A_N^0 and A w.r.t. μ exist and are uniformly bounded in L^2 .

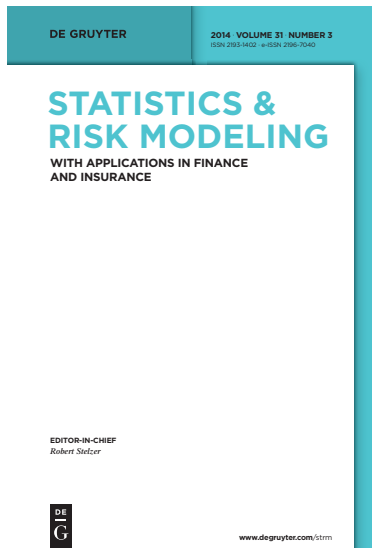
Then the Wigner-Ville spectrum tends pointwise for each $t \in \mathbb{R}$ in mean square to the time-varying spectral density:

$$\int_{\mathbb{R}} |f_N(t, \lambda) - f(t, \lambda)|^2 d\lambda \xrightarrow{N \rightarrow \infty} 0.$$

Thank you very much for your attention!



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